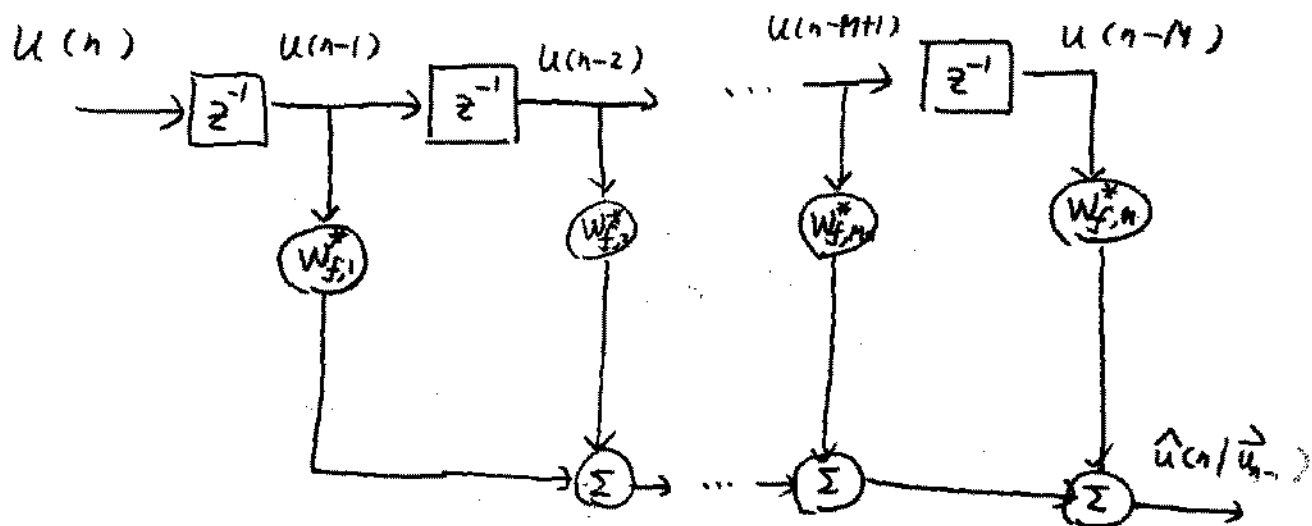
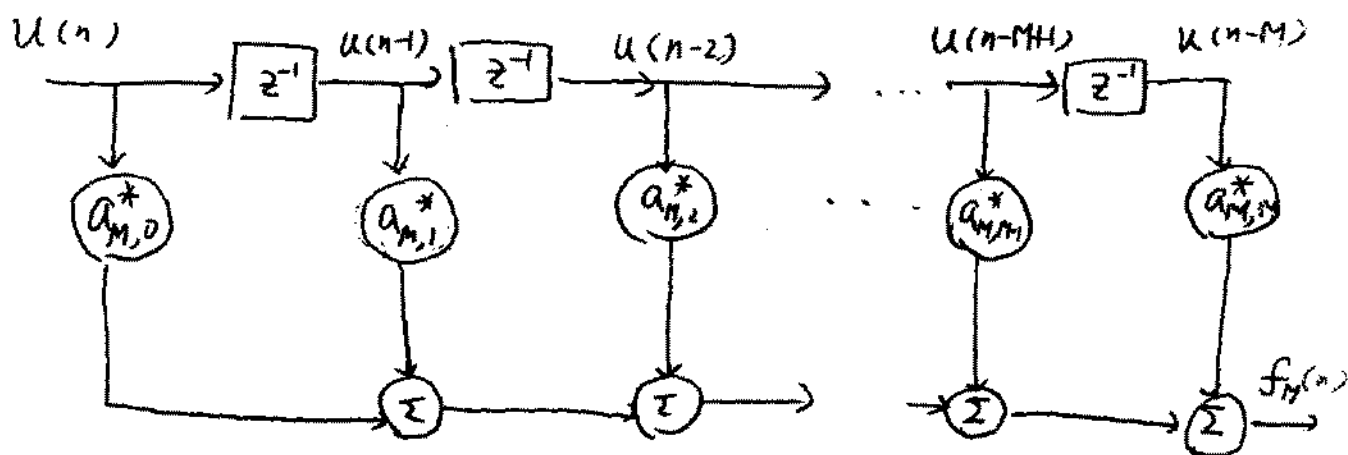


# Chapter 3 Linear Prediction



a linear predictor (one step)



a prediction-error filter

( $a_{m,0} = 1$ )

The predicted value in a linear predictor can be written as

$$\hat{u}(n | \vec{u}_{n-1}) = \sum_{k=1}^M w_{f,k}^* u(n-k)$$

The desired response is  $d(n) = u(n)$ .

Accordingly, the forward prediction error is

$$f_M(n) = u(n) - \hat{u}(n | \vec{u}_{n-1}) \quad \text{and}$$

its mean-square value is

$$P_M = E[|f_M(n)|^2].$$

Based on this linear predictor system, the Wiener-Hopf Equation can be formulated.

Let's denote

$$\vec{w}_f = [w_{f,1}, w_{f,2}, \dots, w_{f,M}]^T$$

$$\vec{u}(n-1) = [u(n-1), u(n-2), \dots, u(n-M)]^T$$

$$\tilde{R} = E[\vec{u}(n-1) \vec{u}^H(n-1)]$$

$$= \begin{bmatrix} r(0) & r(1) & \dots & r(M-1) \\ r^*(1) & r(0) & \dots & r(M-2) \\ \vdots & \vdots & \ddots & \vdots \\ r^*(M-1) & r^*(M-2) & \dots & r(0) \end{bmatrix}$$

The cross-correlation vector

$$\vec{p} = \vec{r} = E [ \vec{u}(n-1) u^*(n) ]$$

$$= \begin{bmatrix} r^*(1) \\ r^*(2) \\ \vdots \\ r^*(M) \end{bmatrix} = \begin{bmatrix} r(-1) \\ r(-2) \\ \vdots \\ r(-M) \end{bmatrix}$$

The Wiener-Hopf Equations based on this linear predictor is

$$\tilde{R} \vec{w}_f = \vec{r}$$

and the minimum mean square prediction-error is

$$P_M = r(0) - \vec{r}^H \vec{w}_f$$

### X: Forward Prediction - Error Filter

If we just add one more connection to the linear predictor, namely, the desired signal  $u(n)$ , then we can construct a prediction-error filter. The output of the

prediction-error filter is just the prediction error sequence  $f_M(n)$ , such that

$$f_M(n) = u(n) - \sum_{k=1}^M w_{f,k}^* u(n-k)$$

$$\Rightarrow f_M(n) = \sum_{k=0}^M a_{M,k}^* u(n-k),$$

$$\text{where } a_{M,k} = \begin{cases} 1, & k=0 \\ -w_{f,k}^*, & k=1, 2, \dots, M \end{cases}$$

X. Augmented Wiener-Hopf Equations for Forward Prediction:

We can combine Eq. (3.9) and (3.10) to construct an augmented Wiener-Hopf Equation such that

$$\begin{bmatrix} r(0) & \vec{r}^H \\ \vec{r} & \vec{R} \end{bmatrix} \begin{bmatrix} 1 \\ -\vec{w}_f \end{bmatrix} = \begin{bmatrix} P_M \\ 0 \end{bmatrix}$$

$$\text{or } \sum_{l=0}^M a_{M,l} r(l-i) = \begin{cases} P_M, & i=0 \\ 0, & i=1, 2, \dots, M \end{cases}$$

$$\text{where } \vec{a}_M = \begin{bmatrix} 1 \\ -\vec{w}_f \end{bmatrix}$$

Example: Compute the  $\vec{a}_M$  for  
a linear predictor of order  $M=1, 2$

Equation (3.14) yields

$$M=1, \begin{bmatrix} r(0) & r(1) \\ r^*(1) & r(0) \end{bmatrix} \begin{bmatrix} a_{1,0} \\ a_{1,1} \end{bmatrix} = \begin{bmatrix} P_1 \\ 0 \end{bmatrix}$$

$$a_{1,0} = \frac{P_1}{\Delta r}, \quad a_{1,1} = -\frac{P_1}{\Delta r} r^*(1),$$

$$\text{where } \Delta r = \det \begin{pmatrix} r(0) & r(1) \\ r^*(1) & r(0) \end{pmatrix} \\ = |r(0) - |r(1)||^2$$

$$P_1 = \frac{\Delta r}{r(0)} \quad \text{since } a_{1,0} = 1$$

$$\Rightarrow a_{1,1} = -\frac{r^*(1)}{r(0)}$$

$$M=2, \begin{bmatrix} r(0) & r(1) & r(2) \\ r^*(1) & r(0) & r(1) \\ r^*(2) & r^*(1) & r(0) \end{bmatrix} \begin{bmatrix} a_{2,0} \\ a_{2,1} \\ a_{2,2} \end{bmatrix} = \begin{bmatrix} P_2 \\ 0 \\ 0 \end{bmatrix}$$

$$a_{2,0} = \frac{P_2}{\Delta r} [r(0) - |r(1)|^2],$$

$$a_{2,1} = -\frac{P_2}{\Delta r} [r^*(1)r(0) - r(1)r^*(2)],$$

$$a_{2,2} = \frac{P_2}{\Delta r} [(r^*(1))^2 - r(0)r^*(2)]$$

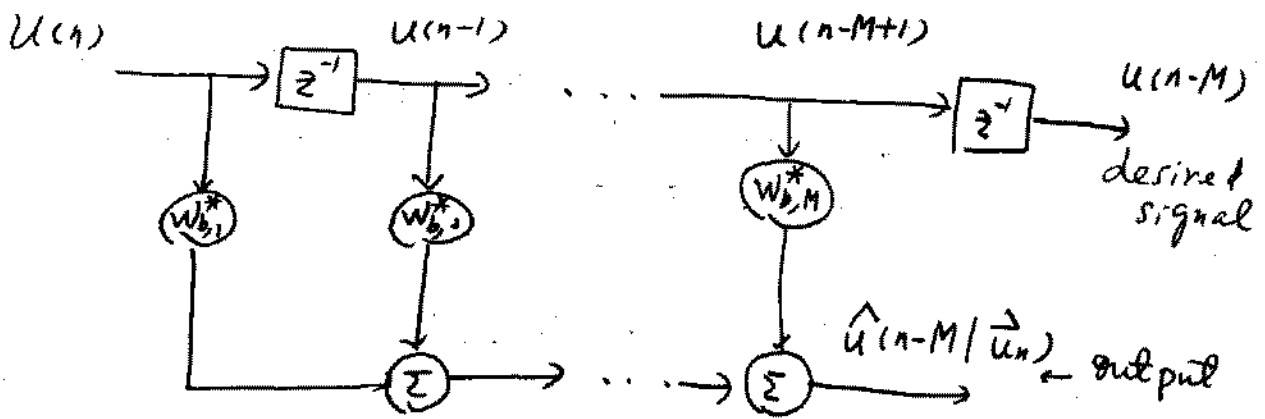
where  $\Delta r = \det \begin{pmatrix} r(0) & r(1) & r(2) \\ r^*(1) & r(0) & r(1) \\ r^*(2) & r^*(1) & r(0) \end{pmatrix}$

Since  $a_{2,0} = 1$ ,  $P_2 = \frac{\Delta r}{r^2(0) - |r(1)|^2}$

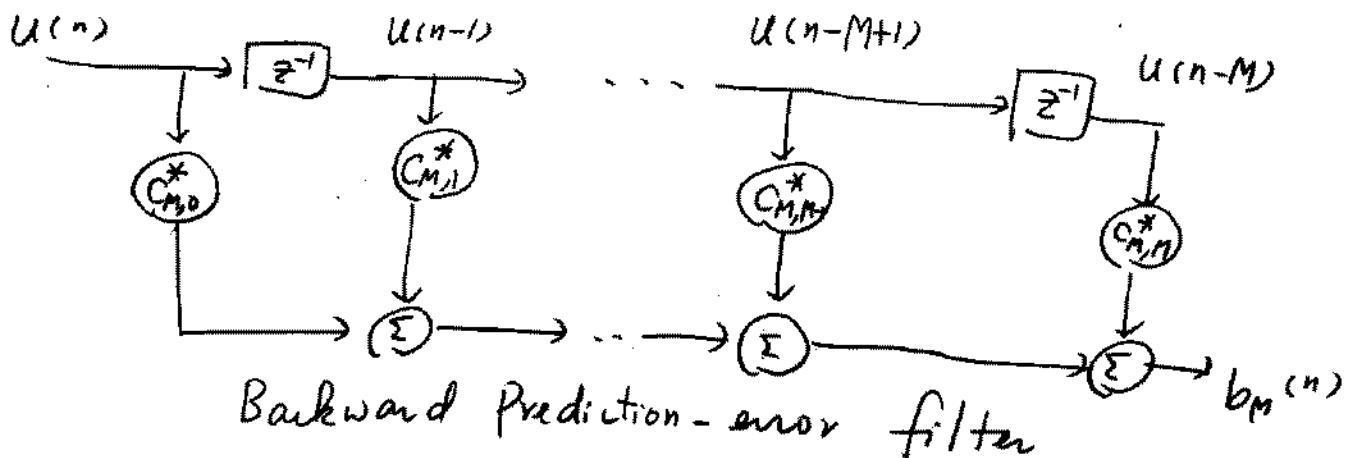
$a_{2,1} = - \frac{r^*(1)r(0) - r(1)r^*(2)}{r^2(0) - |r(1)|^2}$

$a_{2,2} = \frac{(r^*(1))^2 - r(0)r^*(2)}{r^2(0) - |r(1)|^2}$

### 3.2 Backward Linear Prediction



One-step backward predictor



Backward prediction-error filter

The output of a backward predictor can be written as

$$\hat{u}(n-M | \vec{u}_n) = \sum_{k=1}^M w_{b,k}^* u(n-k+1)$$

$$d(n) = u(n-M),$$

and the prediction error is

$$b_M(n) = u(n-M) - \hat{u}(n-M | \vec{u}_n),$$

Thus the mean-square error is

$$P_M = E \{ |b_M(n)|^2 \}, \quad \forall n$$

Let  $\vec{w}_b$  denote the  $M$ -by-1 optimum tap weight vector of the backward predictor in Fig 3-2(a), such that

$$\vec{w}_b = [w_{b,1}, w_{b,2}, \dots, w_{b,M}]^T$$

To find the  $\vec{w}_b$  through Wiener-Hopf Equations,

We need to construct the following stuff:

1. The input vector of this backward predictor is  $\vec{u}(n) = [u(n), u(n-1), \dots, u(n-M+1)]^T$ .

Thus, the correlation matrix is

$$\tilde{R} = E[\vec{u}(n) \vec{u}^H(n)]$$

2. The cross-correlation vector is

$$\vec{p} = \vec{r}^{B*} = E[\vec{u}(n) u^*(n-M)]$$

3. The desired signal energy is

$$E[|d(n)|^2] = E[|u(n-M)|^2] = r(0)$$

Therefore the Wiener-Hopf Equations can be formulated as

$$\tilde{R} \vec{w}_b = \vec{r}^{B*}$$

and the minimum mean-square backward prediction error

is

$$P_M = r(0) - \vec{r}^{B*T} \vec{w}_b$$



# X. Relations between backward and Forward Predictors

Compare Eq. (3.9) and Eq. (3.24), the Wiener-Hopf Equations for forward predictor and backward predictor, respectively,

$$\begin{cases} \tilde{R} \vec{w}_f = \vec{r} \\ \tilde{R} \vec{w}_b = \vec{r}^{B*} \end{cases} \Rightarrow \tilde{R}^T \vec{w}_b^B = \vec{r}^* \\ \Rightarrow \tilde{R}^H \vec{w}_b^{B*} = \vec{r} \\ \Rightarrow \tilde{R} \vec{w}_b^{B*} = \vec{r}$$

Therefore,  $\vec{w}_b^{B*} = \vec{w}_f$

Hence the forward prediction-error in Eq. (3.10)

$$\begin{aligned} P_M &= r(0) - \vec{r}^H \vec{w}_f \\ &= r(0) - \vec{r}^H \vec{w}_b^{B*} \end{aligned}$$

$$P_M^B = P_M = r(0) - \vec{r}^{B^H} \vec{w}_b^*$$

$$(P_M^B)^* = P_M = r(0) - \vec{r}^{B^T} \vec{w}_b \dots \rightarrow \text{Eq. (3.25)}$$

Therefore, we find that the mean-square backward prediction-error has exactly the same value as the mean-square forward prediction-error.  $P_m$  can denote both kinds of prediction errors quantitatively.

### \* Backward Prediction-Error Filter

The backward prediction error  $b_m(n)$  equals

$$b_m(n) = u(n-M) - \sum_{k=1}^M W_{b,k}^* u(n-k+1)$$

If we define the coefficients in such a backward prediction error filter as follows:

$$C_{m,k} = \begin{cases} -W_{b,k+1} & , k=0, 1, \dots, M-1 \\ 1 & , k=M. \end{cases}$$

Hence, 
$$b_m(n) = \sum_{k=0}^M C_{m,k}^* u(n-k)$$

From Eq. (3.28), we have

$$W_{b, n-k+1}^* = W_{f,k} \quad , k=1, 2, \dots, M$$

or 
$$W_{b,k} = W_{f, M-k+1}^* \quad , k=1, 2, \dots, M$$

$$C_{M,k} = \begin{cases} -w_{f,M-k}^* & , k=0, 1, \dots, M-1 \\ 1 & , k=M \end{cases}$$

Compare it with Eq. (3.12), we have

$$C_{M,k} = a_{M,M-k}^* , k=0, 1, 2, \dots, M$$

Hence, the backward prediction error filter output can be written as

$$b_M(n) = \sum_{k=0}^M a_{M,M-k} u(n-k)$$

\* Augmented Wiener-Hopf Equations for Backward Prediction

Combine Eq. (3.24) and Eq. (3.25), we

have

$$\begin{bmatrix} \tilde{R} & \vec{r}^B \\ \vec{r}^{B^T} & r(0) \end{bmatrix} \begin{bmatrix} -\vec{w}_b \\ 1 \end{bmatrix} = \begin{bmatrix} \vec{0} \\ P_M \end{bmatrix}$$

$$\text{or } \sum_{l=0}^M a_{M,M-l}^* r(l-i) = \begin{cases} 0 & , i=0, \dots, M-1 \\ P_M & , i=M \end{cases}$$

### 3.3 Levinson-Durbin Algorithm

We can recursively compute the prediction error filter coefficients from model order  $m = 1, 2, 3, \dots, M$ , where  $M$  is the final order of the filter. Let the  $(m+1)$ -by-1 vector  $\vec{a}_m$  denote the tap-weight vector of a forward prediction error filter of order  $m$ ; then  $\vec{a}_m^{B*}$  denotes the tap-weight vector corresponding to backward prediction error filter coefficients.

The Levinson-Durbin recursion may be stated in one of the following two equivalent ways.

1. The tap-weight vector of a forward prediction-error filter may be order-updated by the equation:

$$\vec{a}_m = \begin{bmatrix} \vec{a}_{m-1} \\ 0 \end{bmatrix} + K_m \begin{bmatrix} 0 \\ \vec{a}_{m-1}^{B*} \end{bmatrix},$$

where  $K_m$  is a constant.

The scalar version of this order update is

$$a_{m,l} = a_{m-1,l} + k_m a_{m-1,m-l}^*, \quad l=0,1,\dots,m,$$

where  $a_{m,l}$  is the  $l^{\text{th}}$  tap-weight of a forward prediction-error filter of order  $m$ , and

$a_{m-1,m-l}^*$  is the  $l^{\text{th}}$  tap-weight of a backward prediction-error filter of order  $m-1$ . Note that

$$a_{m-1,0} = 1, \text{ and } a_{m-1,m} = 0$$

2 The tap-weight vector of a backward prediction-error filter may be order-updated by the equation:

$$\vec{a}_m^* = \begin{bmatrix} 0 \\ \vec{a}_{m-1}^* \end{bmatrix} + k_m^* \begin{bmatrix} \vec{a}_{m-1} \\ 0 \end{bmatrix}.$$

The scalar version of this order update is

$$a_{m,m-l}^* = a_{m-1,m-l}^* + k_m^* a_{m-1,l}, \quad l=0,1,\dots,m,$$

where  $a_{m,m-l}^*$  is the  $l^{\text{th}}$  tap-weight of the backward prediction-error filter of order  $m$ .

The Levinson-Durbin recursion is usually formulated in the context of forward prediction given by Eq. (3.40) or Eq. (3.41).

The formulation of the recursion in the context of backward prediction, in Eq. (3.42) and (3.43) can follow from Eq. (3.40) & (3.41).

To derive Levinson-Durbin algorithm, we proceed in four stages:

1. We premultiply both sides of Eq. (3.40) by  $\tilde{R}_{m+1}$  such that

$$\tilde{R}_{m+1} \vec{a}_m = \begin{bmatrix} P_m \\ 0 \end{bmatrix},$$

where  $P_m$  is the mean-square forward prediction error.

2. Since  $\tilde{R}_{m+1} = \begin{bmatrix} \tilde{R}_m & \vec{r}^{B*} \\ \vec{r}_m^{BT} & r(0) \end{bmatrix}$ ,

$$\begin{aligned} \tilde{R}_{m+1} \begin{bmatrix} \vec{a}_{m-1} \\ 0 \end{bmatrix} &= \begin{bmatrix} \tilde{R}_m & \vec{r}_m^{B*} \\ \vec{r}_m^{BT} & r(0) \end{bmatrix} \begin{bmatrix} \vec{a}_{m-1} \\ 0 \end{bmatrix} \\ &= \begin{bmatrix} \tilde{R}_m \vec{a}_{m-1} \\ \vec{r}_m^{BT} \vec{a}_{m-1} \end{bmatrix} \end{aligned}$$

Iteratively, the set of augmented Wiener-Hopf equations for the forward prediction-error filter of order  $m-1$  is

$$\tilde{R}_m \vec{a}_{m-1} = \begin{bmatrix} P_{m-1} \\ \vec{0}_{m-1} \end{bmatrix},$$

where  $P_{m-1}$  is the mean-square prediction error for this filter of order  $m-1$ . Next, we define a scalar

$$\begin{aligned} \Delta_{m-1} &= \vec{r}_m^T \vec{a}_{m-1} \\ &= \sum_{l=0}^{m-1} r(l-n) a_{m-1, l}. \end{aligned}$$

Substituting Eq. (3.46) and (3.47) into Eq. (3.45), we have

$$\tilde{R}_{m+1} \begin{bmatrix} \vec{a}_{m-1} \\ 0 \end{bmatrix} = \begin{bmatrix} P_{m-1} \\ \vec{0}_{m-1} \\ \Delta_{m-1} \end{bmatrix}$$

3. Since

$$\tilde{R}_{m+1} = \begin{bmatrix} r(0) & \vec{r}_m^H \\ \vec{r}_m & \tilde{R}_m \end{bmatrix},$$

$$\tilde{R}_{m+1} \begin{bmatrix} 0 \\ \vec{B}^* \\ \vec{a}_{m-1} \end{bmatrix} = \begin{bmatrix} r(0) & \vec{r}_m^H \\ \vec{r}_m & \tilde{R}_m \end{bmatrix} \begin{bmatrix} 0 \\ \vec{B}^* \\ \vec{a}_{m-1} \end{bmatrix}$$

$$= \begin{bmatrix} \vec{r}_m^H \vec{B}^* \\ \vec{r}_m \vec{B}^* \\ \tilde{R}_m \vec{a}_{m-1} \end{bmatrix}$$

Consequently,

$$\begin{aligned} \overrightarrow{H} \overrightarrow{a}_{m-1}^* &= \sum_{k=1}^m r^*(-k) a_{m-1, m-k}^* \\ &= \sum_{l=0}^{m-1} r^*(l-m) a_{m-1, l}^* \\ &= \Delta_{m-1}^* \end{aligned}$$

Therefore, the set of augmented Wiener-Hopf equations for the backward prediction-error filter of order  $m-1$  is

$$\widetilde{R}_m \overrightarrow{a}_{m-1}^* = \begin{bmatrix} \overrightarrow{0}_{m-1} \\ P_{m-1} \end{bmatrix}.$$

Substituting Eq. (3.50), (3.51) into Eq. (3.49)

we may write

$$\widetilde{R}_{m+1} \begin{bmatrix} 0 \\ \overrightarrow{a}_{m-1}^* \end{bmatrix} = \begin{bmatrix} \Delta_{m-1}^* \\ \overrightarrow{0}_{m-1} \\ P_{m-1} \end{bmatrix}.$$

4. In summary, combining (3.44), (3.48) and (3.52), we will have

$$\begin{bmatrix} P_m \\ \overrightarrow{0}_m \end{bmatrix} = \begin{bmatrix} P_{m-1} \\ \overrightarrow{0}_{m-1} \\ \Delta_{m-1} \end{bmatrix} + k_m \begin{bmatrix} \Delta_{m-1}^* \\ \overrightarrow{0}_{m-1} \\ P_{m-1} \end{bmatrix}$$



Eg. (3.40) and Eg. (3.53) can provide the complete recursion for order-update from  $m-1$  to  $m$ , in the application of a prediction-error filter.  $P_m = P_{m-1} + K_m \Delta_{m-1}^*$

From Eg. (3.53), two important deductions can be achieved:

$$1. \quad P_m = P_{m-1} + K_m \Delta_{m-1}^*$$

$$2. \quad 0 = \Delta_{m-1} + K_m P_{m-1}$$

$$\Rightarrow K_m = - \frac{\Delta_{m-1}}{P_{m-1}}$$

$$\Rightarrow P_m = P_{m-1} (1 - |K_m|^2)$$

Hence,  $0 \leq P_m \leq P_{m-1}$ ,  $m \geq 1$ ,  $P_0 = r(0)$ ,

deductively,  $P_M = P_0 \prod_{m=1}^M (1 - |K_m|^2)$

\* Interpretation of the Parameters  $K_m$  and  $\Delta_{m-1}$

The parameters  $K_m$ ,  $1 \leq m \leq M$ , in the Levinson-Durbin recursion are called reflection coefficients.

According to Eq. (3.58),

$$|k_m| \leq 1, \quad \forall m$$

According to Eq. (3.41),

$$k_m = a_{m,m}$$

$\Delta_{m-1}$  is interpreted as a cross-correlation between the forward prediction error  $f_{m-1}^{(n)}$  and the delayed backward prediction error  $b_{m-1}^{(n-1)}$ .

$$\Delta_{m-1} = E \left\{ b_{m-1}^{(n-1)} f_{m-1}^{*(n)} \right\},$$

where  $f_{m-1}^{(n)}$  is produced at the output of a forward prediction-error filter of order  $m-1$  in response to the tap inputs  $u(n), u(n-1), \dots, u(n-m+1)$  and  $b_{m-1}^{(n-1)}$  is the output of a backward prediction-error filter of order  $m-1$  in response to the tap inputs  $u(n-1), u(n-2), \dots, u(n-m)$ .

Note that  $f_0^{(n)} = b_0^{(n)} = u(n)$

$$\begin{aligned}\Delta_0 &= E [ b_0^{(n-1)} f_0^{*(n)} ] \\ &= E [ u^{(n-1)} u^{*(n)} ] = r^*(1).\end{aligned}$$

\* Relationship between reflection coefficients and partial correlation coefficients:

The partial correlation (PARCOR) coefficient between the forward prediction error  $f_{m-1}^{(n)}$  and the delayed backward prediction error  $b_{m-1}^{(n-1)}$  is defined as

$$\rho_m = \frac{E [ b_{m-1}^{(n-1)} f_{m-1}^{*(n)} ]}{\sqrt{(E [ |b_{m-1}^{(n-1)}|^2 ] E [ |f_{m-1}^{(n)}|^2 ])}$$

According to the definition, we have

$$|\rho_m| \leq 1, \quad \forall m,$$

$$\text{since } \left| E [ b_{m-1}^{(n-1)} f_{m-1}^{*(n)} ] \right|^2 \leq E [ |b_{m-1}^{(n-1)}|^2 ] \times E [ |f_{m-1}^{(n)}|^2 ]$$

$$\text{From Eq. (3.56) and (3.60),}$$

$$K_m = - \frac{E [ b_{m-1}^{(n-1)} f_{m-1}^{*(n)} ]}{P_{m-1}}$$

$$\text{and } P_{m-1} = E [ |f_{m-1}(n)|^2 ] = E [ |b_{m-1}(n-1)|^2 ]$$

From Eq. (3.61) and Eq. (3.62), we may obtain

$$k_m = -P_m.$$

Example: Use the Levinson-Durbin recursion to determine the corresponding tap weights  $a_{3,1}$ ,  $a_{3,2}$ ,  $a_{3,3}$  and the prediction-error power  $P_3$  for a prediction-error filter of order 3.

$$m = 1, 2, 3$$

1. For a prediction-error filter of order  $m=1$

$$a_{1,0} = 1$$

$$a_{1,1} = k_1, \quad \text{and} \quad P_1 = P_0 (1 - |k_1|^2)$$

2. For a prediction-error filter of order  $m=2$

$$a_{2,0} = 1$$

$$a_{2,1} = k_1 + k_2 k_1^*$$

$$a_{2,2} = k_2 \quad \text{and} \quad P_2 = P_1 (1 - |k_2|^2)$$

3. For a prediction-error filter of order  $m=3$

$$a_{3,0} = 1$$

$$a_{3,1} = a_{2,1} + k_3 k_2^*$$

$$a_{3,2} = k_2 + k_3 a_{2,1}^*$$

$$a_{3,3} = k_3 \quad \text{and} \quad P_3 = P_2 (1 - |k_3|^2)$$

X. Derive the Reflection Coefficients

Define the inner-product of two complex random variables with zero-mean, as

$$\langle x, y \rangle = E \{ x^* y \},$$

$$\text{Therefore } \|x\|^2 = \langle x, x \rangle = E \{ |x|^2 \} = \text{Var}(x).$$

The predicted value can be written as

$$\begin{aligned} \hat{u}_m(n) &= \hat{u}(n | \vec{u}_{n-1}) = \sum_{k=1}^m w_{F,k}^* u(n-k) \\ &= - \sum_{k=1}^m a_{m,k}^* u(n-k) \end{aligned}$$

The mean-square error is

$$\begin{aligned} E_m &= E \{ |u(n) - \hat{u}_m(n)|^2 \} \\ &= \|u(n) - \hat{u}_m(n)\|^2. \end{aligned}$$

$E_m$  will be minimized when the principle of the orthogonality is satisfied, such

$$\text{that } \langle u(n-k), u(n) - \hat{u}_m(n) \rangle = 0, \quad k=1, 2, \dots, m$$

$$\text{Thus, } \langle u(n-k), u(n) + \sum_{l=1}^m a_{m,l}^* u(n-l) \rangle = 0$$

$$\Rightarrow \sum_{l=1}^m a_{m,l}^* \langle u(n-l), u(n-l) \rangle = - \langle u(n-k), u(n) \rangle$$

$$\Rightarrow \sum_{l=1}^m a_{m,l}^* E \{ u^*(n-l) u(n-l) \} = - E \{ u^*(n-l) u(n) \}$$

$$\Rightarrow \sum_{l=1}^m a_{m,l}^* r_u(k-l) = - r_u(k), \quad k=1, 2, \dots, m$$

..... (\*)

$$P_m = \min \{ E_m \} = \langle u(n) - \hat{u}_m(n), u(n) - \hat{u}_m(n) \rangle$$

$$= \langle u(n), u(n) - \hat{u}_m(n) \rangle - \underbrace{\langle \hat{u}_m(n), u(n) - \hat{u}_m(n) \rangle}_{=0}$$

$$= \langle u(n), u(n) \rangle - \langle u(n), \hat{u}_m(n) \rangle$$

$$= \langle u(n), u(n) \rangle + \sum_{k=1}^m a_{m,k}^* \langle u(n), u(n-k) \rangle$$

$$= r_u(0) + \sum_{k=1}^m a_{m,k}^* r_u(-k) \quad \dots (**)$$

Eq (\*) is the well-known Yule - Walker

Equations, or in the matrix form, it

can be stated as follows:

$$\begin{bmatrix} r_u(0) & r_u(-1) & \dots & r_u(-m+1) \\ r_u(1) & r_u(0) & \dots & r_u(-m+2) \\ \vdots & \vdots & \ddots & \vdots \\ r_u(m-1) & r_u(m-2) & \dots & r_u(0) \end{bmatrix} \begin{bmatrix} a_{m,1}^* \\ a_{m,2}^* \\ \vdots \\ a_{m,m}^* \end{bmatrix} = \begin{bmatrix} r_u(1) \\ r_u(2) \\ \vdots \\ r_u(m) \end{bmatrix}$$

$$\tilde{R}_u = \tilde{R}_u^H$$

Alternatively, from Eq. (\*\*), the augmented Yule-Walker Equations can be derived in the matrix form as follows:

$$\begin{bmatrix} r_u(0) & r_u(-1) & \dots & r_u(-m) \\ r_u(1) & r_u(0) & \dots & r_u(-m+1) \\ \vdots & \vdots & \ddots & \vdots \\ r_u(m) & r_u(m-1) & \dots & r_u(0) \end{bmatrix} \begin{bmatrix} 1 \\ a_{m,1}^* \\ \vdots \\ a_{m,m}^* \end{bmatrix} = \begin{bmatrix} p_m \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

Take  $m=1$ ,

$$\hat{u}_1(n) = -a_{1,1} u(n-1)$$

$$\langle u(n-1), u(n) - \hat{u}_1(n) \rangle = 0$$

$$\Rightarrow a_{1,1}^* = - \frac{\langle u(n-1), u(n) \rangle}{\langle u(n-1), u(n-1) \rangle}$$

$$\hat{u}_1(n) = \frac{\langle u(n-1), u(n) \rangle}{\langle u(n-1), u(n-1) \rangle} u(n-1)$$

Denote  $\bar{e}_0^b(n-1) = \frac{u(n-1)}{\|u(n-1)\|}$ , where

$$\|u(n-1)\| = \sqrt{\|u(n-1)\|^2} = \sqrt{\langle u(n-1), u(n-1) \rangle}$$

Therefore, 
$$\hat{u}_1(n) = \frac{\langle u(n-1), u(n) \rangle}{\|u(n-1)\|} \frac{u(n-1)}{\|u(n-1)\|}$$

$$= \langle \bar{e}_0^b(n-1), u(n) \rangle \bar{e}_0^b(n-1).$$

It can be seen that the first-order optimal linear predictor is achieved by projecting  $u(n)$  along the  $u(n-1)$  "direction" where the "unit vector" along the  $u(n-1)$  direction is  $\bar{e}_0^b(n-1)$ .

Now take  $m=2$ ,

$$\hat{u}_2(n) = -a_{2,1}^* u(n-1) - a_{2,2}^* u(n-2)$$

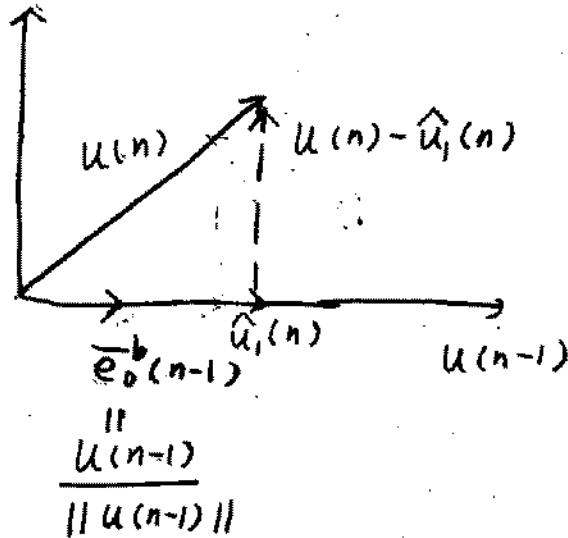
$$= \hat{u}_1(n) + \text{best prediction of } u(n) \text{ based on part of } u(n-2) \text{ in new orthogonal direction.}$$

$$\langle \bar{e}_1^b(n-1), u(n) \rangle \bar{e}_1^b(n-1)$$



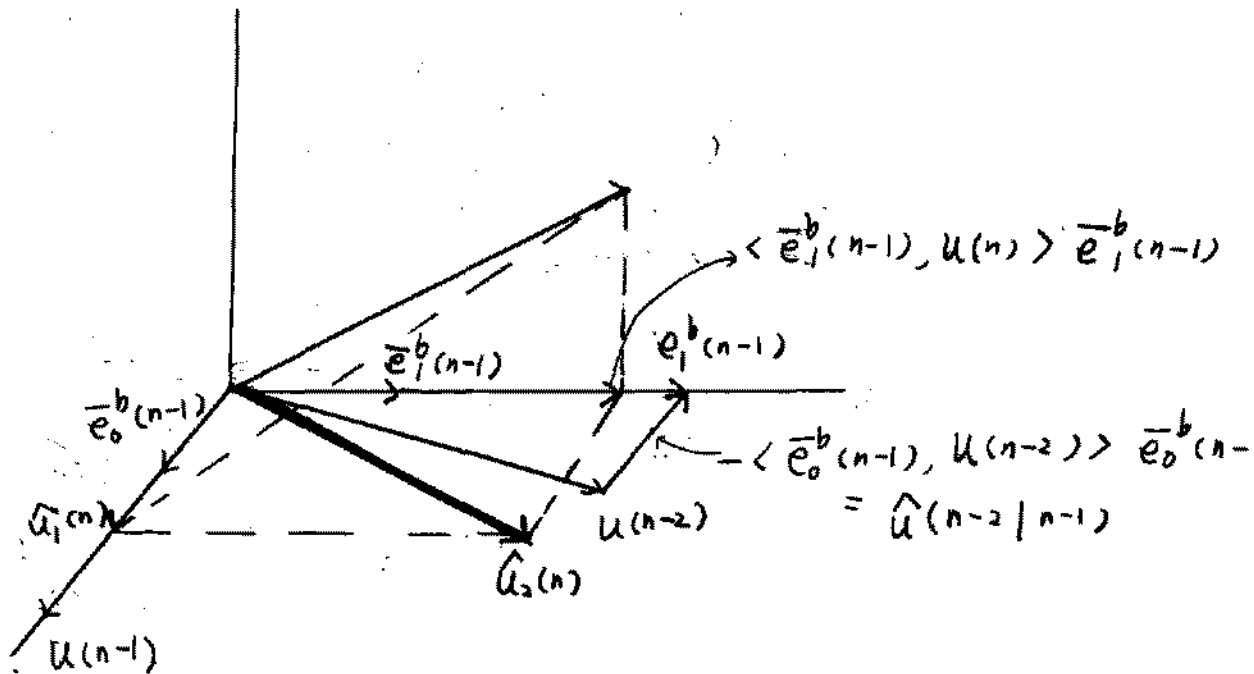
Graphically, the linear prediction can be depicted as follows:

$m=1$



The optimal  $\hat{u}_1(n)$  is the projected vector from  $u(n)$  onto  $\bar{e}_0^b(n-1)$

$m=2$



From the figure above, it clearly shows that  $\hat{u}_2(n)$  consists of two components, namely  $\hat{u}_1(n)$ , and  $\langle \bar{e}_1^b(n-1), u(n) \rangle \bar{e}_1^b(n-1)$ ,

where  $e_1^b(n-1) = u(n-2) - \hat{u}(n-2|n-1)$

and  $\bar{e}_1^b(n-1) = \frac{e_1^b(n-1)}{\|e_1^b(n-1)\|}$

Thus, 
$$\begin{aligned} \hat{u}_2(n) &= \hat{u}_1(n) + \langle \bar{e}_1^b(n-1), u(n) \rangle \bar{e}_1^b(n-1) \\ &= -a_{2,1}^* u(n-1) - a_{2,2}^* u(n-2) \end{aligned}$$

By induction,

$$\hat{u}_m(n) = \hat{u}_{m-1}(n) + \langle \bar{e}_{m-1}^b(n-1), u(n) \rangle \bar{e}_{m-1}^b(n-1),$$

where  $e_{m-1}^b(n-1)$  is the backward prediction error if  $u(n-m)$  is predicted on the basis of  $u(n-(m-1)), u(n-(m-2)), \dots, u(n-1)$ .

$$\hat{u}_m(n) = - \sum_{i=1}^{m-1} a_{m-1,i}^* u(n-i) + \frac{\langle e_{m-1}^b(n-1), u(n) \rangle}{\|e_{m-1}^b(n-1)\|^2} e_{m-1}^b(n-1)$$

$$\text{Let } K_m \equiv \frac{(\langle e_{m-1}^b(n-1), u(n) \rangle)^*}{\|e_{m-1}^b(n-1)\|^2}$$

Note:  $e_{m-1}^b(n-1) \equiv b_{m-1}(n-1)$  as before !!

$$\text{Since } \|b_{m-1}(n-1)\|^2 = \|f_{m-1}(n-1)\|^2 = \|f_{m-1}(n)\|^2,$$

$$\begin{aligned} \text{and } \langle e_{m-1}^b(n-1), u(n) \rangle &= \langle b_{m-1}(n-1), u(n) + \sum_{i=1}^{m-1} a_{m-1,i} u(n-i) \rangle \\ &= \langle b_{m-1}(n-1), f_{m-1}(n) \rangle \quad (\langle b_{m-1}(n-1), u(n-i) \rangle = 0 \\ &\quad, \forall i = 1, 2, \dots, m-1), \end{aligned}$$

$$\text{then } K_m = \frac{(\langle b_{m-1}(n-1), f_{m-1}(n) \rangle)^*}{\|f_{m-1}(n)\| \|b_{m-1}(n-1)\|}$$

$$= \frac{E [ b_{m-1}(n-1) f_{m-1}^*(n) ]}{\|b_{m-1}(n-1)\| \|f_{m-1}(n)\|}$$

$$\sqrt{E [ |b_{m-1}(n-1)|^2 ] E [ |f_{m-1}(n)|^2 ]}$$

as defined before.

The  $m^{\text{th}}$  order predictor becomes

$$\hat{u}_m(n) = - \sum_{i=1}^{m-1} a_{m-1,i}^* u(n-i) + K_m^* b_{m-1}(n-1),$$

$$\text{where } b_{m-1}(n-1) = u(n-m) - \left[ - \sum_{i=0}^{k-2} c_{m-1,i}^* u(n-1-i) \right]$$

Since  $C_{m-1, m-1} = 1$ ,  $C_{m-1, i}^*$

$$\begin{aligned}\hat{u}_m(n) &= - \sum_{i=1}^{m-1} a_{m-1, i}^* u(n-i) - K_m^* \sum_{i=0}^{m-1} C_{m-1, i}^* u(n-1-i) \\ &= - \sum_{i=1}^m a_{m, i}^* u(n-i).\end{aligned}$$

$$\Rightarrow C_{m-1, i} = a_{m-1, m-1-i}^*, \quad i = 0, 1, 2, \dots, m-1$$

According to the principle of the orthogonality,

$$\langle u(n-j), \sum_{i=0}^{m-1} C_{m-1, i}^* u(n-1-i) \rangle = 0, \quad j = 1, 2, \dots, m-1$$

$$\text{or } \sum_{i=0}^{m-1} C_{m-1, i}^* r_u(j-i-1) = 0.$$

Let  $l = m-1-i$ , we obtain

$$\sum_{l=0}^{m-1} C_{m-1, m-1-l}^* r_u(j-m+l) = 0$$

Take the complex-conjugate, we have

$$\sum_{l=0}^{m-1} C_{m-1, m-1-l} r_u(m-j-l) = 0, \quad j = 1, 2, \dots, m-1.$$

$$\Rightarrow \sum_{l=0}^{m-1} C_{m-1, m-1-l} r_u(m-l) = 0, \quad m = 1, 2, \dots, m-1$$

$$\Rightarrow a_{m-1, l}^* = C_{m-1, m-1-l}, \quad l = 0, 1, 2, \dots, m-1$$

$$\begin{aligned} \hat{u}_m(n) &= - \sum_{i=1}^{m-1} a_{m-1,i}^* u(n-i) - k_m^* \left( u(n-m) + \sum_{i=0}^{m-2} a_{m-1,m-1-i} u(n-1-i) \right) \\ &= - \sum_{i=1}^{m-1} (a_{m-1,i}^* + k_m^* a_{m-1,m-i}^*) u(n-i) \\ &\quad - k_m^* u(n-m). \end{aligned}$$

Since  $\hat{u}_m(n) = - \sum_{i=1}^m a_{m,i}^* u(n-i)$ ,

therefore  $a_{m,i}^* = a_{m-1,i}^* + k_m^* a_{m-1,m-i}^*$   
 or  $a_{m,i} = \begin{cases} a_{m-1,i} + k_m a_{m-1,m-i}^*, & i=1,2,\dots,m \\ k_m, & i=m \end{cases}$

### 3.4. Properties of Prediction-error Filters

Property 1. The corresponding set of reflection coefficients can be obtained based on the set of autocorrelation function  $r(1), \dots, r(M)$ ,

such that  $k_m = - \frac{1}{P_{m-1}} \sum_{k=0}^{m-1} a_{m-1,k} r(k-m)$ .

Proof: From Eq. (3.47),

$$\Delta_{m-1} = \sum_{l=0}^{m-1} r(l-m) a_{m-1, l}$$

From Eq. (3.55),

$$\Delta_{m-1} = -k_m P_{m-1}$$

Hence 
$$\sum_{l=0}^{m-1} a_{m-1, l} r(l-m) = -k_m P_{m-1}$$

Since  $r(m) = r^*(-m)$  and  $a_{m-1, 0} = 1$ ,

we get 
$$k_m = -\frac{1}{P_{m-1}} \sum_{l=0}^{m-1} a_{m-1, l} r(l-m),$$

$$m = 1, 2, \dots, M$$

Example: Provided  $P_0$ ,  $k_1$ ,  $k_2$  and  $k_3$ ,  
compute  $r(0)$ ,  $r(1)$ ,  $r(2)$  and  $r(3)$ .

Solution:

$$m=1 \Rightarrow k_1 = -\frac{1}{P_0} a_{0,0}^* r(1)$$

$$\Rightarrow r(1) = -P_0 k_1^*$$

where  $P_0 = r(0)$ .

$$m=2 \Rightarrow K_2 = -\frac{1}{P_1} (a_{1,0}^* r(2) + a_{1,1}^* r(1))$$

$$\text{where } P_1 = P_0 (1 - |k_1|^2)$$

$$\Rightarrow r(2) = -P_1 K_2^* - r(1) k_1^*$$

$$m=3 \Rightarrow K_3 = -\frac{1}{P_2} (a_{2,0}^* r(3) + a_{2,1}^* r(2) + a_{2,2}^* r(1))$$

$$\text{where } P_2 = P_1 (1 - |k_2|^2)$$

$$\Rightarrow r(3) = -P_2 K_3^* - a_{2,1}^* r(2) - K_2^* r(1),$$

$$\begin{aligned} \text{where } a_{2,1} &= a_{1,1} + k_2 a_{1,0}^* \\ &= k_1 + k_2 k_1^* \end{aligned}$$

Property 2. The transfer function of a forward prediction - error filter can be stated as follows:

$$H_{f,m}(z) = H_{f,m-1}(z) + K_m^* z^{-1} H_{b,m-1}(z)$$

$$\text{where } H_{f,m}(z) = \sum_{k=0}^m a_{m,k}^* z^{-k} \text{ and}$$

$$H_{b,m-1}(z) = \sum_{k=0}^{m-1} c_{m-1,k}^* z^{-k}$$

Property 3. A forward prediction-error filter is minimum phase.

Proof: The mean-square forward prediction error

$$\rho = E \{ |u(n) - \hat{u}(n)|^2 \}$$

$$= E \left\{ \left| \sum_{k=0}^m a_{m,k} u(n-k) \right|^2 \right\}$$

The minimum prediction error power can

$$\text{be written as } P_m = \int_{-\frac{1}{2}}^{\frac{1}{2}} S_{uu}(f) \left| 1 + \sum_{k=1}^m a_{m,k} e^{-j2\pi f k} \right|^2 df$$

$$\text{Let } S_a(e^{j2\pi f}) = 1 + \sum_{k=1}^m a_{m,k} e^{-j2\pi f k}$$

$$S_a(z) = \prod_{j=1}^m (1 - z_j z^{-1}), \text{ where } z_j$$

is the  $j^{\text{th}}$  zero of  $S_a(z)$

If some  $z_i$  is outside the unit circle,

$$S_a(z) = (1 - z_i z^{-1}) \underbrace{\prod_{\substack{j=1 \\ j \neq i}}^m (1 - z_j z^{-1})}_{S'_a(z)}$$

where  $|z_i| > 1$

$$\text{Then } P_m = \int_{-\frac{1}{2}}^{\frac{1}{2}} \left| 1 - z_i e^{-j2\pi f} \right|^2 |S'_a(e^{j2\pi f})|^2$$

$$\times S_{uu}(f) df$$

$$> \int_{-\frac{1}{2}}^{\frac{1}{2}} \left| 1 - \frac{1}{z_i^*} e^{-j2\pi f} \right|^2 |S'_a(e^{j2\pi f})|^2 S_{uu}(f) df$$



$\Rightarrow P_m$  is not minimum prediction error!  
 Therefore no such  $z_i$ , where  $|z_i| > 1$  can exist!  
 Since  $\tilde{R}_{m+1}$  is positive definite, no  $z_j$  can be on the unit circle either.  
 $\Rightarrow z_j$  must be within the unit-circle  $\Rightarrow$  minimum phase.

Property 4. A backward prediction-error filter is maximum-phase.

Proof: 
$$H_{f,m}(z) = \sum_{k=0}^m a_{m,k} z^{-k}$$

$$H_{f,m}^* \left( \frac{1}{z^*} \right) = z^m \sum_{k=0}^m a_{m,m-k} z^{-k}$$

$$= z^m H_{b,m}(z)$$

The zeros in  $H_{f,m}(z)$  are assumed to be  $z_i$ ,  $i = 1, 2, \dots, m$

Hence  $H_{f,m}(z_i) = 0$  or  $H_{f,m}^*(z_i^*) = 0$

Then,  $H_{b,m} \left( \frac{1}{z_i^*} \right) = H_{f,m}^*(z_i^*) = 0$

$\therefore H_{b,m}(z)$  has  $m$  zeros as  $\frac{1}{z_i^*}$ ,  
 $i = 1, 2, \dots, m$ .

Since  $H_{f,m}(z)$  has minimum-phase zeros  $z_i$ ,  $H_{b,m}(z)$  will have all maximum phase zeros  $\frac{1}{z_i^*}$ .

Property 5. A forward prediction error filter is a whitening filter.

Proof: By orthogonality principle, it can be easily shown that

$$E \{ f_m^*(n) f_m(n-k) \} = \begin{cases} P_m, & k=0 \\ 0, & k \neq 0 \end{cases}$$

### 3.6 Autoregressive Modeling of a Stationary random Process

We may view the operation of prediction-error filtering applied to a random process  $u(n)$  as one of the autoregressive output response and the white process  $f_m(n)$

is the excitation input. The AR modeling can be implemented in two stages:

1. Given  $u(n), u(n-1), \dots, u(n-M)$ , then estimate the optimal  $a_{m,k}$ ,  $k=0, 1, \dots, M$ , such that  $f_m(n)$  is white. (Analysis)
2. Excite the AR filter with the determined  $a_{m,k}$ ,  $k=0, 1, 2, \dots, M$ , obtained from stage 1, with arbitrary new white process  $v(n)$  to replace  $f_m(n)$  to generate the artificial AR outputs. (Synthesis)

### 3.7. Cholesky Factorization

Since the optimal solution based on the Wiener-Hopf Equations would involve with the inverse of a large matrix  $\tilde{R}$ , an efficient computational algorithm can be stated here as the Cholesky Factorization.

The backward prediction-error sequences can be expressed as:

$$b_0(n) = u(n)$$

$$b_1(n) = a_{1,1} u(n) + a_{1,0} u(n-1)$$

$$\vdots$$

$$\underline{b_M(n)} = \underline{a_{M,M}} u(n) + \underline{a_{M,M-1}} u(n-1) + \dots + \underline{a_{M,0}} u(n-M)$$

Hence,

$$\underline{\vec{b}}(n) = \underline{\tilde{L}} \underline{\vec{u}}(n)$$

where  $\underline{\vec{u}}(n) = [u(n), u(n-1), \dots, u(n-M)]^T$ ,

$$\underline{\vec{b}}(n) = [b_0(n), b_1(n), \dots, b_M(n)]^T,$$

and

$$\underline{\tilde{L}} = \begin{bmatrix} 1 & 0 & \dots & 0 \\ a_{1,1} & 1 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ a_{M,M} & a_{M,M-1} & \dots & 1 \end{bmatrix}$$

Orthogonality of Backward Prediction Errors

It can be easily derived such that

$$E \{ b_m(n) b_i^*(n) \} = \begin{cases} P_m, & i=m \\ 0, & i \neq m \end{cases}$$

-X Factorization of the inverse Correlation Matrix

$$E \left[ \begin{matrix} \vec{b}^{(n)} \\ \vec{b}^{(n)H} \end{matrix} \right] = E \left[ \begin{matrix} \tilde{L} \vec{u}^{(n)} \\ \vec{u}^{(n)H} \tilde{L}^H \end{matrix} \right]$$

$$= \tilde{L} E \left[ \begin{matrix} \vec{u}^{(n)} \\ \vec{u}^{(n)H} \end{matrix} \right] \tilde{L}^H$$

Denote  $\tilde{D} = E \left[ \begin{matrix} \vec{b}^{(n)} \\ \vec{b}^{(n)H} \end{matrix} \right]$

$$= \tilde{L} \tilde{R}^H = \begin{bmatrix} P_0 & 0 & \dots & 0 \\ 0 & P_1 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & \dots & 0 & P_M \end{bmatrix}$$

"diag  $\{P_0, P_1, \dots, P_M\}$ "